
DOCUMENT
DE TRAVAIL
N° 286

**ARE THE EFFECTS OF FISCAL CHANGES DIFFERENT IN TIMES OF
CRISIS AND NON-CRISIS? THE FRENCH CASE**

C. Bouthevillain and G. Dufrenot

July 2010



**ARE THE EFFECTS OF FISCAL CHANGES DIFFERENT IN TIMES OF
CRISIS AND NON-CRISIS? THE FRENCH CASE**

C. Bouthevillain and G. Dufrénot

July 2010

Les Documents de travail reflètent les idées personnelles de leurs auteurs et n'expriment pas nécessairement la position de la Banque de France. Ce document est disponible sur le site internet de la Banque de France « www.banque-france.fr ».

Working Papers reflect the opinions of the authors and do not necessarily express the views of the Banque de France. This document is available on the Banque de France Website “www.banque-france.fr”.

Are the effects of fiscal changes different in times of crisis and non-crisis? The French Case

Carine Bouthevillain¹

Gilles Dufrénot²

The authors wish to thank Valérie Mignon for her fruitful comments and suggestions as a discussant in a seminar at Banque de France as well as the participants in the Banca d'Italia workshop on “Fiscal Policy: What Lessons from the Crisis?” held in Perugia, 25-27 March 2010. The views expressed in the paper are those of the authors and do not reflect those of the institutions to which they belong.

¹ Banque de France, Direction de la Conjoncture et de la Prévision Macroéconomique, Service des Finances Publiques, 39 rue Croix des Petits Champs, 75001, Paris, Tel : +33 1 42 92 42 92, Email: carine.bouthevillain@banque-France.fr

² Banque de France, CEPII and DEFI, Université de la Méditerranée, Centre de Recherche en Développement Économique et Finance Internationale, 14 Avenue Jules Ferry, 13621, Aix-en-Provence, Cedex, Tel : +33 4 42 91 48 34, Email : gilles.dufrenot@univmed.fr

Abstract

This paper shows that the impact of changes in budgetary variables on major macroeconomic variables varies in sign and magnitude in times of crisis and non-crisis in France. We find that these nonlinearities are both frequent (as they exist on all behaviors analyzed: real GDP, private consumption, business investment and private employment) and significant. For this, we estimate time-varying probability Markov-switching models (TVPMS) in order to take into account two budgetary regimes, on the one side periods of severe recessions or depressions (crises), and, on the other side “normal” periods (expansions or moderate recessions). These two regimes are identified endogenously, so that we do not need to preliminary separate episodes of huge contractions and expansions of the business cycle. Further, we are able to identify the variables influencing the probability of a switch between regimes. Searching for nonlinear fiscal impacts in the form of regime-switching effects, we assume temporary variations in the budgetary variables, both on the revenue side (taxes on consumption, on firm’s profit, lump sum transfers) and on the expenditure side (traditional public boosts of aggregate demand, transfers, and subsidies). Our results show that if one considers the aggregate GDP, public expenditure has a stronger impact during crisis and the expenditure multiplier is greater than the tax multiplier. Also, when households are sensitive to the unemployment situation, tax cuts do not increase consumption spending, while transfers are playing a significant role. On the firms side, our results show that direct taxes changes induce a (stimulus) effect in the investment rate only during non-crisis periods. A rise in subsidies has a negative influence during crises. Finally, the estimates suggest that employment policies should be asymmetric: fiscal measures aiming at reducing unit labor costs could be efficient in good times, while an increase in public employment is preferable during crisis.

Keywords : Markov-switching models, fiscal policy, crisis

JEL classification : C51, E62, H50.

Résumé

Cette étude montre que, en France, l'impact de changements budgétaires sur les principales variables macro-économiques varie à la fois en termes de signe et d'ampleur selon que l'on est ou pas en période de crise. Nous montrons que ces non-linéarités sont à la fois fréquentes (elles existent pour tous les comportements analysés : le PIB réel, la consommation privée, l'investissement des entreprises et l'emploi privé) et significatives. Pour cela, un modèle de type Markov-Switching avec des probabilités variables dans le temps (TVPMS) est estimé, afin de considérer deux régimes budgétaires, d'une part un régime de graves récessions ou de dépressions (crises), et, d'autre part les périodes «normales» (expansion ou récession modérée). Ces deux régimes sont identifiés de façon endogène, de sorte que nous n'avons pas besoin d'identifier préalablement les épisodes de contraction et d'expansion majeurs du cycle économique. En outre, nous identifions les variables qui influencent la probabilité de passer d'un régime à l'autre. Pour trouver des non-linéarités dans les impacts budgétaires sous la forme de changement de régime, nous considérons des variations temporaires dans les variables budgétaires, tant du côté des recettes (impôts sur la consommation, sur le bénéfice entreprise, transferts forfaitaires) que du côté des dépenses (outils traditionnels de relance de la demande globale, transferts et subventions). Nos résultats montrent que si l'on considère le PIB total, les dépenses publiques ont un impact plus fort pendant une crise et le multiplicateur de dépenses est supérieur au multiplicateur de recettes. Par ailleurs, lorsque les ménages sont sensibles à la situation du chômage, les réductions d'impôt n'augmentent pas les dépenses de consommation, alors que les transferts jouent un rôle important. Du côté des entreprises, nos résultats montrent que les baisses d'impôts directs produisent un stimulus sur le taux d'investissement uniquement pendant les périodes de non-crise. Une augmentation des subventions a une influence négative en période de crise. Enfin, les estimations indiquent que les politiques de l'emploi devraient être asymétriques : les mesures fiscales visant à réduire les coûts unitaires du travail pourraient être efficaces dans les périodes favorables, alors qu'une augmentation de l'emploi public est préférable en cas de crise.

Mots clés : modèle Markov-switching, multiplicateur budgétaire, politique budgétaire, crise

Codes JEL : C51, E62, H50.

1.- Goal of the paper and motivation

This paper shows that the impact of changes in budgetary variables on real GDP, investment, consumption and employment varies in sign and magnitude in times of crisis and non-crisis. To this end, a regime-switching process is embedded in standard macroeconomic equations in order to take into account different budgetary regimes. Our purpose is threefold.

First, we aim at reconsidering the non-monotonic effects of fiscal policy over the business cycle by distinguishing, on the one side periods of severe recessions or depressions (crises), and, on the other side “normal” periods (expansions or moderate recessions). For illustration purpose, we consider the French case, since our study can help in judging the quantitative impact of the fiscal package (“plan de relance”) undertaken by the French fiscal authorities in 2008. Indeed, the latter has revived conflicting views about its effectiveness. On the one hand, the estimates provided by the government plead in favor of strong Keynesian effects which are supposed to help the economy to lean against the crisis. On the other hand, less favorable outcome of fiscal multiplier estimates elsewhere in the World has led to an increasing sentiment of “failure” of Keynesian belief regarding the positive impact of budgetary variables on the real GDP (For studies concerning other countries, see Cogan et al. (2010), Trabandt and Uhlig (2009), Uhlig (2009)). These views are however not necessarily contradictory, since both Keynesian and non-Keynesian effects may be observed at different times.

Secondly, we consider the nonlinear response of a variety of fiscal measures targeted to private consumption, business investment, private employment, in addition to the real GDP. Indeed, non-monotonic responses to fiscal changes are likely to be more precisely estimated if we consider the components of the GDP but not only the real GDP itself. The reason is that, the nonlinear response of the GDP to fiscal changes most of the time can be explained by the private-sector’s behavior (because any policy modifies market confidence, expectations among the public about future outcome and accordingly the agents’ decisions).

Thirdly, and more importantly, we are searching for nonlinear fiscal impacts in the form of regime-switching effects. The following paragraphs explain our motivations for doing this.

Doubts about the successfulness of the recent massive fiscal interventions in the world rely on the recognition that there are fiscal regimes and that the latter alternate in a stochastic way. Regime-switching approaches to modeling fiscal policy have been an important aspect of the theoretical literature in endogenous growth models. Fiscal policy regimes have been identified as Keynesian or Ricardian regimes, low debt-output or high debt-output regimes, passive and active regimes, etc³.

³In a pioneering paper, Sutherland (1997) shows that high public debt during times of crisis may reverse the effects of fiscal policy in an exogenous stochastic growth model. Davig (2004) derives regime-switching macroeconomic equilibria from an endogenous growth model in which agents face a signal extraction problem

The key idea is that the economy is unstable – and unpredictable - in terms of its reaction to budgetary changes due to two features. The first feature is the time-varying nature of fiscal policy reaction functions. Fiscal interventions vary over time in terms of magnitude and in terms of the instrument used (tax or spending) according to governments' policy objectives, to the macroeconomic environment and to the state of public finances (fiscal space)⁴. Since changes in fiscal policy switch in stance and nature due to political and economic circumstances, they are better understood by relating them to different regimes. The second feature is the changing nature of the cyclical response to fiscal changes because agents' reaction to budgetary policy depends upon elements that are not under the direct control of the governments themselves (liquidity constraints, adjustment costs, leverage effects, Barro-Ricardo effects, credit market imperfection, etc⁵). The combination of the two features implies that the effects of fiscal policy on the real activity are characterized by stochastic changes over time in the multipliers. Researchers usually model these changes in two manners.

A common approach, mainly empirical, consists in providing evidence of asymmetric effects of fiscal changes on the economy between regimes that are defined according to a prior belief by the researcher: expansion and recession phases in the business cycle, times of fiscal contractions and fiscal expansions, regimes of active and passive budgetary rules, large and persistent or small and non-persistent fiscal impulses, times of binding liquidity constraints and "good" times, etc.. The models contain dummy variables that capture structural breaks or threshold functions allowing for a dependence of fiscal multipliers to the level of an exogenous variable (for instance public debt ratio)⁶.

An alternative approach, mainly theoretical, relies on the simulations of general equilibrium-based models in which fiscal rules (determining spending, taxes, or debt) are governed by a two-state Markov chain variable and agents make a probabilistic inference regarding the future rule and state of the economy to take their decisions. These models are based on the assumption of asymmetric information between governments and the private sector (firms and households). The latter thus use Bayesian procedures to learn the regime generating the expected future variables on which they base their investment and consumption decisions (debt-output ratio, tax, or spending)⁷.

on forthcoming fiscal policies. Minea and Villieu (2008) propose an endogenous growth model à la Barro which exhibit a regime-switching effect of fiscal deficits on economic growth, depending on public debt ratio.

⁴ There are examples in the literature of regime-switching tests of fiscal behaviors (see Favero and Monacelli (2005), Thams (2006), Claeys (2008)).

⁵ For a survey of the different factors implying more or less effective fiscal policy, see Beetsma (2008).

⁶ For typical examples, we refer the reader to Perotti (1999), Giavazzi et al. (2000), Giavazzi et al.(2005), Minea and Villieu (2008), Tagkalakis (2008).

⁷ See Dotsey (1994), Ruge-Murcia (1995), Dotsey and Mao (1997), Davig (2004).

This paper adopts the second approach. Since, we search to differentiate the budgetary effects on the macroeconomic variables between times of crisis and non-crisis, we can assume that the root cause of the differing fiscal effects is the high uncertainty facing the public and private sectors. Crises appear occasionally, suddenly, with no specific regularity; they are characterized by huge depressions that make them different from standard business cycle troughs. Further, their duration is not predictable. For governments, in such a context, fixed-regime rules are hard to implement, fiscal policy requires more flexibility and is likely to be characterized by more or less frequent switches rather than by stability. Their decisions to cut taxes, or increase spending, are therefore influenced by their forecasts of the future state of the economy. Their belief can be represented by probabilities. For the private sector, profit- and consumption-maximizing decisions are influenced by fiscal policy and, as shown in the aforementioned papers, agents solve a signal extraction problem when the information on both the state of the economy and fiscal policy is incomplete and asymmetric. These decisions are well described in a probabilistic framework involving Markov-switching variables.

Though we adopt the Markov-switching framework to study the non-monotonic effects of fiscal policy in times of crisis and non-crisis, our approach differs from those of the previous papers in the literature in the sense that it is not theoretical. Instead, we add to the previous literature by considering econometric models. Simulations derived from micro-founded models provide us with qualitative features, which need to be completed with quantitative measures. We thus consider a set of reduced-form equations that can be derived from the Markov-switching general equilibrium models mentioned in footnote 1, and, we estimate them.

We estimate time-varying probability Markov-switching models (TVPMS) to see whether the effects of fiscal policy on the real economy vary in France between times of crisis and non-crisis. These two regimes are identified endogenously, so that we do not need to preliminary separate episodes of huge contractions and expansions of the business cycle. Further, we are able to identify the variables influencing the probability of a switch between regimes. We assume temporary variations in the budgetary variables and focus our attention on the effectiveness of fiscal measures at stimulating aggregate demand and output in the short-run. This seems realistic as during exceptionally severe crises governments' fiscal measures consist of temporary interventions and are centered on Keynesian demand management and fine-tuning of the business cycle. Prices and the exchange rate are thus assumed to be fixed and fiscal changes only cause aggregate demand variables to fluctuate.

We examine the effects of various types of taxes (on consumption, on firms' profits, lump sum transfers) and various targets for government spending (traditional boosts of aggregate demand, subsidies). A common wisdom for modeling the effects of shocks is to compute impulse response functions after "shocking" the non-systematic component (innovations) of the budgetary variables. Another way to proceed, used in this paper, consists in introducing a stochastic process in the coefficients of estimated equations where the parameters are regime-dependent and where the manner in which regime shifts occur is specified by a probability distribution function defining the

probability of transition from either regime to another. In this type of models, changes in the budgetary variables are considered as intra-regime shocks. For instance, a typical question is: what is the short-run impact of a 1% change in government spending on the output if the likelihood that the economy is in a crisis regime is high? In this alternative approach, the uncertainty is not due to the fact that shocks are unanticipated, but to the fact that even when they are expected, the current state of the economy is not observed ex-ante.

Finally, we do not distinguish between the discretionary and non-discretionary changes in the fiscal variables, but consider the effects of changes in the budgetary variables taken as a whole. Indeed, the effectiveness of fiscal changes depends upon both discretionary stimulus and the size of automatic stabilizers.

The paper is structured as follows. Section 2 presents the estimated equations. Section 3 discusses the econometric methodology of time-varying transition Markov-switching models. Section 4 presents the results, while Section 5 elaborates on some policy implications. Finally, Section 6 concludes.

2. – Benchmark equations

In this section we lay out the equations that are estimated to study the nonlinear effects of budgetary policies between times of crisis and non-crisis. We consider four endogenous variables: first, private GDP; second, private consumption; third, business investment and fourth, employment. Each variable is fairly standard in macroeconomic models, the difference here being that we want to see which circumstances are most likely to give rise to a non-monotonic response of these variables to budgetary changes, be they positive (expansionary fiscal policy) or negative (consolidations).

Our reduced-form equations are linearised versions of the solutions derived from the theoretical set-ups mentioned in footnote 1, which introduce Markov-switching stochastic processes in micro-founded models of the economic growth. One difference is however the nature of the regimes that we consider. Since the theoretical models often focus on fiscal regimes, the regimes are defined accordingly. For instance, Davig (2004) distinguishes between a low debt-output regime and a high debt-output regime. Dotsey (1994) makes a difference between a low tax regime and a high tax regime. Here, the regimes are those of crisis and non-crisis. We neither impose any ex-ante restriction about what is called a “crisis”, nor on the years when the latter occurs. We simply keep in mind that, usually, a crisis is characterized, first by severe depressions (drop of the output and of the main components of aggregate demand) and secondly by shifts in key macroeconomic and policy variables (public debt-ratio, taxes and spending, output-gap, credit demand, etc). Since, we do not know ex-ante the regime (“crisis” or “non-crisis”) generating the observed changes in the real GDP, consumption, investment or employment, we assume that the agents make a probabilistic inference on their occurrence, regarding the state of some key macroeconomic and policy variables (called

transition variables) which reflect the “circumstances” under which the economy is likely or not likely to switch from either regime to the other.

Since the Markov-switching models are defined under the assumptions that all our variables are stationary, we consider the first-differences of the exogenous/endogenous variables and the transition variables alike⁸. Besides, since our intention is to study the regime-switching effects of fiscal policy, in our benchmark equations, we assume that the switching between regimes is only driven by the fiscal variables (in addition to the lagged terms of the endogenous variables). Our equations include lags on the endogenous variables in order to capture costs of adjustments or partial adjustment dynamic behaviors.

Consumption is modeled as a function with the real income as the main control variable. It is augmented with habit-persistence behavior (lags) and fiscal variables such as direct taxes and transfers. This equation can be derived from a theoretical model where households aim at maximizing a utility function upon consumption and labor, for given values of their revenues, taxes and transfers. We assume that labor supply is inelastic to the real wages in a context of high unemployment rate.

We further consider business investment and private employment equations that are assumed to be derived from profit maximization subject to a Cobb Douglass type production function with the inputs of capital and labor. The maximization behavior leads to linear functions in which capital and labor demands depend upon total demand (the real GDP) and the costs of factors (respectively the real long-run interest rate and the unit labor cost). Since the employment and investment equations are estimated separately, we neglect the cross-equation restrictions imposed by the theory. The equations are completed with fiscal variables that influence the costs of production and the accelerator in the case of the investment.

Finally, we estimate an equation of the real private GDP that corresponds to the reduced form obtained after combining the different behavioral equations and the different identities in a standard general equilibrium model.

Real private GDP

From standard arguments, changes in real private GDP y_t , are explained by control variables, namely the variations in the degree of openness, $open_t$, the real short-term interest rate, i_t , and budgetary variables F_t :

$$\Delta y_t = \varphi_1(s_t) + \lambda(s_t)\Delta y_{t-1} + \varphi_2\Delta open_{t-i} + \varphi_3\Delta i_{t-j} + \varphi_4(s_t)\Delta F_t + \sigma_y\xi_t \quad (1)$$

⁸ We applied unit root tests to our series, in a preliminary step, and concluded in favor of a rejection of the null of no unit root when they were in level. To avoid too many tables, the results are not reported but available upon request to authors.

i, j (in indexes) are lags selected according to information criteria (AIC/BIC) and specification tests on the residuals (serial correlation and remaining nonlinearities). Δ denotes first-differences. ΔF_t is a vector of contemporaneous and lagged changes of the budgetary variables. ξ_t is a stochastic disturbance with a variance σ_y . In our regressions, the best estimates (according to criteria described in the next section) were obtained when the growth rate or public debt or debt-GDP ratio were chosen as the transition variables.

Real private consumption

We estimate the following equation, whose dependent variable is the first-difference of private real consumption:

$$\Delta c_t = \rho_0(s_t) + \rho_1(s_t)\Delta c_{t-1} + \rho_2\Delta w_t + \rho_3(s_t)\Delta transf_t + \sigma_c\vartheta_t \quad (2)$$

ϑ_t is an error term with a variance σ_c^2 . w_t is a vector of contemporaneous and lagged values of households' real disposable income. Nominal income is defined as the sum of wages, households' other revenues (including financial revenues) and individual enterprises' EBITDA (earnings before interests, taxes, depreciation and amortization). $transf_t$ is a vector of contemporaneous and lagged values of transfers. Nominal transfers are positive if they are paid to households (for instance, social payments) and negative if they are paid by households (for instance contribution to social security). The "best" transition variable in our regressions is changes in unemployment.

Real business investment

We consider changes in firms' real investment rate, $\Delta invest_t$, as a function of contemporaneous and lagged changes in real GDP, Δy_t , in the real long-term interest rate, ΔR_t (both variables are in the vector of control variables X_t) and the following fiscal variables enter in the vector F_t : changes in corporate taxes, variations in subsidies and government spending. j, k and l are lags determined by information criteria. The equation is the following:

$$\Delta invest_t = \theta_0(s_t) + \theta_1(s_t)invest_{t-1} + \theta_2\Delta X_t + \theta_4(s_t)\Delta F_t + \sigma_{inv}\omega_t \quad (3)$$

ω_t is an error term with a variance σ_{inv} , ΔX_t is the vector of contemporaneous and lagged changes of the control variables and ΔF_t is the vector of contemporaneous and lagged changes of the budgetary variables. The transition variable is the output-gap (a proxy for the capacity utilization level).

Employment

Changes in private employment, ΔL_t , depend on the growth rate of current and past real GDP, represented by the vector $(\Delta RGDP_t)$, on the variations of the unit labor costs (ratio of unit wages to labor productivity $(\Delta wages_{t-i}/PROD_{t-i})$). Adjustment costs are modeled by the lagged endogenous

variable and we also consider public investment, $INVEST_{t-j}$. i and j are lags. Fiscal policy is assumed to influence two explanatory variables: on the one hand, the unit labor cost varies with, for instance, the employers' contribution to social security or taxes on labor demand; on the other hand, public investment is strongly correlated with government current expenditure and can be considered as an element of public demand. The transition variable is the variations of the output-gap. The equation is the following:

$$\Delta L_t = \tilde{\varphi}_0(s_t) + \tilde{\varphi}_1(s_t)\Delta L_{t-1} + \tilde{\varphi}_2\Delta RGDP_t + \tilde{\varphi}_3(s_t) \left[\Delta \left(\frac{wages_{t-i}}{PROD_{t-i}} \right) \right] + \tilde{\varphi}_4(s_t)\Delta INVEST_{t-j} + \sigma_L \vartheta_t \quad (4)$$

ϑ_t is the error term with a variance σ_L^2 .

3. - Time-varying probability Markov-switching models

3.1.- Definition

We consider an endogenous variable y_t which "visits" two regimes, one corresponding to times of crisis and the other to "normal times". The occurrence of a regime is referred by a variable s_t that takes two values: 1 if the observed regime is 1 and 2 if it is regime 2⁹. We assume that $t=1, \dots, T$.

The observation of either regime 1 or 2 at time t depends upon the regimes visited by the endogenous variable during the previous periods, that is s_t is conditioned by $s_{t-1}, s_{t-2}, \dots, s_{t-k}$. At any time $\tau < t$, the regime that will be observed at time t is unknown with certainty. We thus introduce a probability P of occurrence of s_t given the past regime. Assuming, for purpose of simplicity, that s_t is a first-order Markov-switching process, we define

$$P\{s_t/s_{t-1}, s_{t-2}, \dots, s_{t-k}\} = P\{s_t / s_{t-1}\}. \quad (5)$$

We further assume that the transition from one regime to the other depends upon a set of "transition" variables described by a vector z_t so that

$$P\{s_t / s_{t-1}\} = P\{s_t/s_{t-1}, z_t\}. \quad (6)$$

The relation between z_t and s_t is given by:

$$s_t = \begin{cases} 1, & \text{if } \eta_t < a(s_{t-1}) + z_t' b(s_{t-1}) \\ 2, & \text{if } \eta_t \geq a(s_{t-1}) + z_t' b(s_{t-1}) \end{cases} \quad (7)$$

⁹ We do not discuss here the question as whether the number of states is equal to or different from 2. This is an assumption in our case. However, several methodologies have been proposed to deal with the testing of the number of states to which we refer the interested reader (see, among others, Hamilton (1991), Hansen (1992), Garcia (1998)).

where η_t is distributed as a Φ law. We accordingly define the transition probabilities as follows:

$$\begin{cases} P\{s_t = 1/s_{t-1} = j, z_t\} = p_{1j}(z_t) = \Phi(a_j + z_t' b_j) \\ P\{s_t = 2/s_{t-1} = j, z_t\} = p_{2j}(z_t) = 1 - \Phi(a_j + z_t' b_j) \end{cases} \quad (8)$$

where Φ is either the standard Logistic or Normal cumulative distribution function¹⁰.

Since the dynamics of the endogenous variable is assumed to be regime-dependent, then any influence of explanatory variables, represented by a vector x_t , may differ across regimes. We thus consider the following relationship:

$$y_t = \begin{cases} x_t' \beta_1(s_t) + \sigma_1(s_t) \varepsilon_t, & \text{with a probability } p_1(z_t) \\ x_t' \beta_2(s_t) + \sigma_2(s_t) \varepsilon_t, & \text{with a probability } p_2(z_t) \end{cases} \quad (9)$$

where $\varepsilon_t \sim N(0,1)$. $p_1(z_t)$ and $p_2(z_t)$ are the posterior (or unconditional probabilities) of regimes 1 and 2. The usual probabilistic properties for the ergodicity and the invertibility of (9) applies if we assume that y_t, x_t and z_t are covariance-stationary.

The above model can be generalized to a higher number of states (see Kim et al. (2008)) and encompasses several classes of Markov-switching models previously proposed in the literature (Goldfeld and Quandt (1973), Diebold et al. (1994), Filardo (1994), Hamilton (1989)).

3.2.- Estimation and methodological issues

The above model is estimated via maximum likelihood (henceforth ML) with relative minor modifications to the nonlinear iterative filter proposed by Hamilton (1989). We define the following vectors: $\Omega_t = (X_t, Z_t)$ the vector of observations of x and z up to period t ; $\xi_t = (y_t, y_{t-1}, \dots, y_1)$; $\theta = (\beta_1, \sigma_1, a_1, b_1, \beta_2, \sigma_2, a_2, b_2)$.

The conditional likelihood function of the observed data ξ_t is defined as

$$L(\theta) = \prod_{t=1}^T f(y_t/\Omega_t, \xi_{t-1}; \theta) \quad (10)$$

where

$$\begin{aligned} f(y_t/\Omega_t, \xi_{t-1}; \theta) = & \sum_i \sum_j f(y_t/s_t = i, s_{t-1} = j, \Omega_t, \xi_{t-1}; \theta) \\ & \times P(s_t = i, s_{t-1} = j/\Omega_t, \xi_{t-1}; \theta) \end{aligned} \quad (11)$$

The weighting probability in (11) is computed recursively by applying Bayes's rule:

¹⁰ Any functional form of the transition probabilities that maps the transition variables into the unit interval would be a valid choice for a well-defined log-likelihood function: logistic or Probit family of functional forms, Cauchy integral, piecewise continuously differentiable variables. The choice of a Logistic and Normal law is common wisdom in the applied literature.

$$\begin{aligned}
& P(s_t = i, s_{t-1} = j / \Omega_t, \xi_{t-1}; \theta) \\
& = P(s_t = i / s_{t-1} = j, z_t) P(s_{t-1} = j / \Omega_t, \xi_{t-1}; \theta) \\
& = P_{ij}(z_t) P(s_{t-1} = j / \Omega_t, \xi_{t-1}; \theta)
\end{aligned} \tag{12}$$

We also have

$$\begin{aligned}
& P(s_t = i / \Omega_{t+1}, \xi_t; \theta) = P(s_t = i / \Omega_t, \xi_t; \theta) \\
& \frac{1}{f(y_t / \Omega_t, \xi_{t-1}; \theta)} \sum_j f(y_t / s_t = i, s_{t-1} = j, \Omega_t, \xi_{t-1}; \theta) \\
& \quad \times P(s_t = i, s_{t-1} = j / \Omega_t, \xi_{t-1}; \theta)
\end{aligned} \tag{13}$$

To complete the recursion defined by the equations (11) and (12), we need the regime-dependent conditional density functions

$$f(y_t / s_t = 1, s_{t-1} = j, \Omega_t, \xi_{t-1}; \theta) = \frac{\phi\left(\frac{y_t - x_t' \beta_1}{\sigma_1}\right) \Phi(a_j + z_t' b_j)}{\sigma_1 P_{1j}(z_t)} \tag{14a}$$

$$f(y_t / s_t = 2, s_{t-1} = j, \Omega_t, \xi_{t-1}; \theta) = \frac{\phi\left(\frac{y_t - x_t' \beta_2}{\sigma_2}\right) \Phi(a_j + z_t' b_j)}{\sigma_2 P_{2j}(z_t)} \tag{14b}$$

The parameters of Equations (8) and (9) are thus jointly estimated with ML methods for mixtures of Gaussian distributions. As compared with other estimators (for instance, the EM algorithm or the Gibbs sampler¹¹), the ML estimator has the advantage of computational ease. As shown by Kiefer (1978), if the errors are distributed as a normal law, then the ML yields consistent and asymptotically efficient estimates. Further, the inverse of the matrix of second partial derivatives of the likelihood function at the true parameter values is a consistent estimate of the asymptotic variance-covariance matrix of the parameter values.

The influence of z_t on P_{1j} and P_{2j} gives information about the way the transition variables influence the probability of being in either regime or another. For instance, if regime 1 is the crisis regime, a positive (resp. negative value) of b_1 (resp b_2) implies that the transition variable raises the probability of evolving in a time of crisis.

The optimal combination of the lags on the control and transition variables is determined by computing information criteria (Akaike and Schwarz) for each estimated model. To assess the fit of the estimated models to the data, we apply Ljung-Box tests to the expected standardized residuals as well as tests of remaining nonlinearities (Hinich and Patterson (1989)'s Portmanteau bispectrum test and Tsay (1996)'s test). The expected residuals are the weighted residuals with the weights equal to the probability of observing regimes 1 and 2 at each date.

¹¹ See Diebold et al. (1994) and Filardo and Gordon (1993).

4.- Data and results

We apply the model to France. Data are quarterly, span the years from 1970 to 2009, and are taken from the OECD database. Time series for public finance variables were available at a yearly frequency and were interpolated to get quarterly observations. In order to avoid spurious dynamics stemming from the interpolation method, we simply estimate a “trend” between two observations. Except when their values are negative, the data are transformed into logarithm. Further, we take the first-differences to cope with non-stationarity (unit root tests, available upon request to the authors, showed that the data contain a stochastic trend). The appendix contains a description of data and definition of the variables.

We select the best estimated equations according to the information criteria (AIC/BIC), the inexistence of serial correlation in the residuals, the likelihood ratio test for TVPMS (the null hypothesis is constant probabilities). For each model, the initial values are those of a linear regression of the endogenous variables on the control and fiscal variables.

To avoid endogeneity biases due to the correlation between the endogenous variables budgetary variables, we use a two-step approach by first estimating a VAR system in level composed of the variables of the different equations¹². Then, in a second step, we consider the forecasted in-sample values of the explanatory variables to apply the TVPMS model. As the second stage is linear in the variables, the two-step approach is applicable.

4.1.- Real private GDP equation

Table 1a through 1c report the estimates obtained for the GDP equation. All the variables are expressed in real terms (they are deflated by the GDP deflator). The transition variable is the fourth-order moving average of the differentiated logarithmic real debt or debt ratio. The model detects two regimes corresponding respectively to periods of crisis (huge troughs in the real GDP cycle) and “normal periods” (expansions or moderate recessions). The model improves over a simple constant probability model à la Hamilton. Indeed, the likelihood ratio test for TVPMS is significant (the p-value lies under 5%), thereby indicating a rejection of the hypothesis of constant transition probabilities. Figures 1 through 3 report the smoothed posterior probabilities of either regime 1 or 2 and we see that the smoothed probabilities approach 1 for the two years corresponding to the troughs of 1992/1993 and 2009. The model thus dichotomizes between a regime of crisis (regime 2) and a regime of non-crisis (regime 1). This is shown in Table 1a by the intercepts that are respectively negative (-0.013) and positive (0.005) in each regime. These intercepts capture the average GDP growth within each regime.

¹² By applying a Johansen test, we checked that the variables were cointegrated in levels.

INSERT FIGURES 1 THROUGH 3 HERE

In Table 1a, evidence of an asymmetric effect of public expenditure is assessed by two different coefficients for regimes 1 and 2. Although both regimes are Keynesian (the estimated coefficients are positive), the impact of changes in government spending on the real GDP is higher when the economy is in crisis (regime 2) with a differing effect of 13% (in comparison with regime 1). An increase in public expenditure is therefore efficient to boost real GDP growth, in both times of crisis and non-crisis even though the impact is superior during crises. The control variables have the expected signs. A higher degree of openness increases the real private GDP, while a rise in the real short-term interest rate reduces it (though the latter does not appear to be statistically significant).

Changes in public debt across a year appeared to be the best transition variable (according to various criteria: residual tests, AIC/ BIC criteria, remaining nonlinearities tests). This variable provides information on the fact that any increase in the stock of debt may be interpreted by the private sector as a phenomenon paving the way to possible solvability and sustainability problems in the future. This can decrease the “performance” of the expenditure multiplier if the expectations yield Ricardian behaviors (people save the additional revenues stemming from the new expenditure to pay the future taxes). In terms of our econometric model, the probability of being in a “strong” multiplier regime (regime 2) should decrease if Ricardian behaviors are at work. In this case, we would expect a negative sign of the coefficient b_2 (and a positive sign of b_1) in equation (8). As is seen in Table3, this is *not* the case.

On the other hand, a positive growth rate of the stock of debt implies a higher volume of expenditure, which could raise the magnitude of the impact on the real GDP if private investment and consumption fully and positively respond to public spending. In this case, we would instead expect a positive value of the coefficient b_2 and a negative value of b_1 (with at least one of both coefficients being statistically significant). To say it another way, a rise in public debt lowers the probability of being in regime 1, a regime in which public expenditure have the less significant impact on real GDP growth. This is the case here, as evidenced by the estimated coefficients. This would mean that, in France, there seems not to be Ricardian effects associated with an increase in the stock of debt. Such anti-Keynesian effects do not appear when we consider the aggregate real GDP. Instead, during the crisis regimes, increasing debt provides a fiscal space that reinforces the effects of government spending on the real GDP.

INSERT TABLE 1A ABOUT HERE

We further consider the difference between the growth rate of government expenditure and that of potential output, as an explanatory fiscal variable (instead of changes in government spending). The idea is that in the medium term, a large part of public expenditure is supposed to change according to potential GDP growth (in this case expenditure ratio to GDP remains constant). Then, a positive

difference reflects a discretionary budgetary expansion, while a negative difference means an active fiscal consolidation.

Table 1b lists the estimates corresponding to this case. Again regimes 1 and 2 are respectively classified into “non-crisis” and “crisis” phases (see also Figure 2). However, the above conclusions change. Indeed, if we consider the effects of discretionary public spending (and not the combined effects of the discretionary and automatic stabilizers components of government expenditure, as is the case in Table 1a) the estimates suggest a non-monotonic effect of government spending with a positive and significant impact of the real GDP during crises, but no impact during non-crisis periods. An explanation may be the following. During crises, liquidity constraints are important and reinforce the impact of government expenditure on the activity. During non-crisis periods, crowding-out effects (a decreased in private investment due to the fact that government spending use up resources that would be available otherwise to the private sector) moderate the positive impact of the discretionary policy (this is confirmed further by the estimation of our investment equation). Another point that appears in Table 1b is that the delays of transmission of public spending to the activity differ whether we consider only the discretionary component of public spending or public expenditure as a whole. In the first case, the transmission to the activity takes a longer time (the optimal lag for the government spending variable is 5 in Table 1b, while it is 2 in Table 1a).

INSERT TABLE 1B ABOUT HERE

Table 1c shows estimates when the budgetary variable is the ratio of government revenues to GDP. The estimates are consistent with two different regimes characterized respectively by huge falls of real GDP (regime 1) and increases or moderate decreases in real GDP (regime 2) – see also Figure 3. The fiscal effect on GDP is statistically null in the second regime, but negative and statistically significant in the first. Accordingly, raising fiscal revenues is not harmful for the economy in times of “non-crisis”, but may reduce production when the economy evolves in a crisis phase. Conversely, tax cuts can help to exit from a depression. How can we explain the asymmetric effect of tax revenues of the real GDP? Tax revenues affect production indirectly through their impact on aggregate expenditure (because they involve changes in disposable income, the cost of factors, wealth, etc). If the government reduces taxes with the goal of warding off a huge recession or depression, the increased disposable income of the private sector will be partly consumed and partly saved depending upon the propensity to consume, invest, import, etc. If these propensities are higher in times of crisis as compared with times of non-crisis (due for instance to liquidity constraints), then we can expect a stronger impact when the economy is evolving in a huge trough of the business cycle.

The control variables have the expected signs, respectively positive for the degree of openness and negative for the real short-run interest rate (though the latter does not carry a statistically significant sign).

INSERT TABLE 1C ABOUT HERE

4.2.- Real private consumption

Table 2 shows the results for real private consumption when the unemployment rate is the transition variable. The theoretical literature points that, among the circumstances in which consumption may respond non-monotonically to fiscal variables, the uncertainty about the state of the economy is an important factor. On the one side, during depressions liquidity constraints bind thereby implying a more effective budgetary policy than during non-crisis periods because wealth effects play less importantly. On the other side, if the economy is composed of a non-negligible fraction of *non* rule-of-thumb consumers (who smooth their consumption according to the discounted value of their future income) then a budgetary expansion could have a non-Keynesian effect on consumption during huge downturns of the economic activity.

INSERT TABLE 2 ABOUT HERE

In France, we *do not* find any non-monotonic effect of fiscal policy on real private consumption between regimes of strong falls in consumption (crisis) and regimes of non-crisis, be the instruments taxes on income or social security transfers. The regimes identified by the model are plotted in Figures 4a and 4b. We see that the first regime is described as one in which consumption evolves in a trough. As indicated by the coefficients in Table 2, income taxes have no effects on real private consumption while the effects of transfers appear to be symmetric as we find a coefficient of quite similar size for both crisis and non-crisis regimes (around 0.14). Only the contributions to social security are associated with an asymmetric impact on consumption with a negative outcome only during times of crises.

The probability of being in a crisis regime increases with the unemployment rate, as expected (b_1 carries a positive sign). Finally, the real disposable income positively influences private consumption.

To summarize, only spending increases in the form of transfers to households raise the real private consumption (we have a Keynesian outcome for this variable), but the impact is symmetric. The finding that taxes have no significant effects on consumption can be interpreted with reference to several approaches of the economic literature. For instance, if we consider the effect of tax cuts, we can think that, during crises, there are non-Keynesian effects due to precautionary saving (as the unemployment rate increases) that offset the positive effect on consumption. The size of precautionary saving may be more or less important depending upon whether households face strong liquidity constraints or not. Tax cuts are “consumed” if households are highly constrained (a situation observed during crises) and saved otherwise. This can explain why we obtain a negative sign for the income tax variable in the regime of crisis (-0.0068), but a positive one for the non-crisis regime (0.044). It is possible that the unemployment rate (which is our transition variable) determines whether households take or not their decision of consumption expenditure (in response to a tax decrease or increase) regarding their perceived permanent disposable income. When the unemployment is growing moderately or is decreasing (non-crisis regime), households are more inclined to smooth consumption in comparison with a situation in which the unemployment rate is

increasing fast (as is observed in a crisis regime). In the latter case, consumption is constrained by their current income and this reduces the effect of precautionary saving.

INSERT FIGURES 4A AND 4B ABOUT HERE

4.3.- Business investment

The estimates for business investment are reported in Table 3 and the smoothed posterior probabilities of being in either a regime of sustained increases in investment (regime 1) or in a regime of prolonged decreased (regime 2) are shown in figures 5a and 5b. As seen in Figure 5a, the probability of the second regime “jumps” to 1 around some years that are generally considered as being times of crisis or important recessions : second oil price shock years, the year 1983 which was characterized by a restrictive budgetary policy, 1993, 2001-2002 and, as expected, 2009. Conversely, in Figure 5a, we observe that the probability of being in regime 1 increases during the times when business evolves on an ascending trend.

INSERT FIGURE 5A AND FIGURE 5B AROUND HERE

The outcome of cuts in corporate taxes is an increase in investment in times of booming investment (regime 1). We indeed obtain a statistically significant coefficient of -0.08. Conversely, to mitigate an investment downturn, the instrument of direct tax does not prove efficient as the coefficient is statistically not different from 0 at the 5% level of significance. One reason may be that, during the phases of a depressed activity, firms are more sensitive to demand-side variables than to fiscal discretionary measures.

Our results also point to a significant crowding-out effect of government spending on business investment only in times of booming investment (regime 1)(the coefficient is around -0.39). As is known from theory, there are several channels at play here. The reduction in business investment may occur because the spending is accompanied by a tax increase. As, we have just seen, any increase in corporate taxes does not have a significant impact on firms’ investment behavior periods of booming investment (regime 1). Another mechanism is a reduction in private investment following a higher government borrowing. We tried to use the debt ratio as a transition variable to see whether this variable influences the reaction of business investment to government spending, but it appears not to be conclusive in explaining the asymmetries observed in the data. Crowding-out effects appears to be moderate during recessions or depressions (here non-significant in regime 2) because government spending expands the demand facing the private sector (through the multiplier) thereby implying an accelerator effect that is strong when firms suffers from unused capacities (stronger during the crises than during expansions). In the regression, we can see that the

coefficients related to the impact of the real GDP are big in comparison to the others (the coefficients of lagged GDP terms sum to 1.56).

Government subsidies also appear to have an asymmetric impact on business investment with possible non-Keynesian effects in the second regime (crisis). The subsidies do not influence private investment during expansion phases – the coefficient is not statistically significant in regime 1- but reduce it during recessions. One explanation can be that, during recessions, in addition to reducing capacities, firms also proceed to other internal adjustments (for instance, they deleverage to clean up their balance sheets or reduce their debts).

Turning our attention to the impact of the control variables, we see that the real GDP has an expected positive influence, while the real long-run interest rate acts negatively.

The diagnostic tests show that, while there are no residual correlations (the p-value of the Ljung-Box statistics are above 5%), the residuals still contain remaining nonlinearities (both the Hinich and Tsay tests reject the null hypothesis of linearity). Accordingly, the investment behavior may obey to other type of nonlinearities.¹³

INSERT TABLE 3 ABOUT HERE

4.4.- Private employment

We now consider the asymmetric impact of unit labor costs and public investment on private employment. The economic theory provides us with some explanations about the reason why the effects of changes in these variables on private employment are regime-dependent and depends upon the state of the economy as reflected by the level of demand (huge recessions or not).

The different ways the enterprises respond to the increase in public demand can lead to asymmetric reactions of private employment to changes in public investment. On the one hand, if, in response to higher total demand, they extend their existing capacity level with the same technology, this leads an upward shift of labor demand. On the other hand, if the additional investments incorporate labor saving technology, this leads negative employment effects. The positive demand-side effects are, in general, the result of higher expected profits. These are likely to occur during crises if, for instance, firms are facing strong liquidity constraints. Conversely, enterprises can choose to take advantages of the productivity gains associated with booms or expansions and accordingly to save labor.

A fall in unit labor costs (measured by the ratio of unit wages to total productivity) can lead to an increase in employment as long as labor demand is sensitive to these costs. Changes in the unit labor costs mirror either institutional changes in the labor market structure (for instance labor market

¹³ For instance, since this variable is more volatile than the other components of total demand, nonlinearities may exist in the variance. However, considering these nonlinearities here would make the model cumbersome to estimate).

deregulation inducing less labor market rigidities) or fiscal policy changes (for instance, a fall in taxes on the labor factor of production reducing the unit wages). A decline in labor costs is expected to improve firms' competitiveness and thereby labor demand. But, labor demand may not change at all if enterprisers are very sensitive to the demand for output (this happens during times of crisis). Even, a fall in the unit labor costs that originates from an increase in productivity has two sided effects. On the one hand, it lowers the unit labor cost. On the other hand, it may have a negative impact on employment if the increase in productivity is not matched by an increase in the level of demand (in such a situation firms choose to resort to labor saving).

In our estimations, reported in Table 4, we retrieve these different effects. But before commenting our results, we briefly indicate what the results show. If we look at the raw series of private employment (Figure 6a), we see that its dynamics is described as fluctuations evolving along a quadratic trend (or a piecewise linear trend with a break around 1994). We removed this trend by passing the data through a filter (Hodrick-Prescott)¹⁴. Looking at the residual series, in Figure 6b, we see that the amplitude of the private employment cycle increases from the mid-eighties, which leads to a TVPMS model that is non-stationary if estimated (the highest amplitude reflects the fact that the duration of recessions and expansions becomes higher from the mid-nineties onwards). We thus take the first-difference of the filtered series, which means that we are modeling the *steepness of changes* in private employment. In Figure 6c, we superpose the two graphs representing the steepness in private employment changes and changes in the real GDP (first lag). We see that there is a great concordance between the troughs and peaks of both variables, thereby reflecting that the state of the demand is an important factor in the decision to lay-off or to hire more or less importantly in the private sector. We accordingly choose the output-gap as our transition variable in the TVPMS model.

INSERT TABLE 4 ABOUT HERE

INSERT FIGURES 6A THROUGH 6D ABOUT HERE

Figure 6d shows that the posterior probability of being in regime 1 is around 1 for the years that are usually identified as years of crises (for instance the 2009 crisis, 1992-1993 or the years following the two oil price shocks of the seventies and eighties). The estimated autoregressive coefficients, in Table 4, accord well with the fact that episodes of huge negative variations in private employment occur much more rarely than those of moderate diminutions or increases. The latter are more frequently observed so that the corresponding state is very persistent.

In the second regime (non-crisis), a decrease in unit labor costs comes along with an increase in private employment (the negative coefficient, -0.12, indicates a negative relationship between the

¹⁴ The results are unchanged if we consider the residuals of a regression in which the explanatory variables are an intercept and two linear trends. However, the use of this approach would yield the criticism that the break date may be endogenous.

two variables), while during times of crisis a fall in unit labor costs is accompanied by decreases in labor demand (as illustrated by the positive coefficient, 0.03). This findings reflects the inability of downward pressure in the cost of labor to stimulate employment if, at the same time, total demand is decreasing importantly as is the case in times of crisis.

The results also show asymmetric effects as regards the impact of public investment. We find that any increase results in higher employment in times of crisis (the coefficient carries a positive sign of 0.01), but a fall in non-crisis times. It may be the case that public investment appears as “manna” to firms when they face outlet constraints and that they trade-off between labor and productivity in non-crisis times.

As regards the other coefficients, we find that the higher the value of the output-gap (the higher the value of actual production above potential output), the less likely the probability of evolving in the first regime (crisis), which accords with the fact that in the latter firms have many unused capacities (b_1 is negative and statistically significant). The coefficient of the real GDP carries the expected positive sign. The Chi-squared test yields a significant p-value at the 10% statistical level, which indicates that the TVPMS model has better performance to reproduce the dynamics of the endogenous variable than a standard Hamilton, or linear, model . There are no autocorrelations in the estimated residuals, but we see that, just as with investment, there are remaining nonlinearities that suggest some difficulties in modeling the demand for factors.

5.- Policy implications

We now discuss some policy implications of our findings. The French recovery plan in the aftermath of the crisis was driven by some reductions in taxes and by a raise of public expenditure. Government spending increases accounts for the lion’s share of this plan, so that we can say that it was mainly spending-oriented. However, beyond the crisis fiscal sustainability objectives will come back into the policymakers’ agenda. This raises several important questions. Do we have reason to doubt the effectiveness of the standard Keynesian policy, as suggested by some economists.? Do we observe nonlinear effects in the response of real GDP, private consumption, investment and employment to changes in taxes or spending (for instance, is the response of the economy likely to be weaker or higher during the crisis to a fiscal stimulus, than during the exit-crisis period)? To what extend will it be possible to conciliate both objectives of achieving fiscal sustainability and sustaining economic growth beyond the crisis?

These questions are important because France should begin a process of major fiscal adjustment (4 points off the cyclically-adjusted balance over a period of 3 years are enrolled in the revised stability program presented in January 2010). A central issue is whether such adjustment may have a relatively limited negative effect on growth. Our model can help to shed new light on this point by showing two distinct regimes associated with multipliers with different value or even sign.

What can we conclude about the effects of budgetary variables on the real GDP in France? First, there is evidence of asymmetric effects for both the multiplier of government expenditure and the fiscal multiplier, with differing effects during the phases of crisis and non-crisis. The following table summarizes our findings regarding the impact of the budgetary variables.

Table 5. Effects of budgetary variables (times of crisis and non-crisis)

	Non-crisis regime		Crisis regime	
	Coefficient	T-stat	Coefficient	T-stat
	Impact on real GDP			
Δ government expenditure	0.25	2.75	0.37	3.94
Δ government expenditure – Δ real potential GDP	0.05	1.01	0.296	2.45
Δ Public revenue	-0.044	-1.032	-0.257	-2.19
	Private employment			
Δ unit labor cost	-0.122	-8.38	0.033	4.19
Δ public investment	-0.028	-3.64	0.016	3.25
	Business investment			
Δ corporate taxes	-0.08	-2.21	0.022	0.76
Δ subsidies	0.048	1.27	-0.17	-3.04
Δ government spending	-0.394	-2.42	-0.357	-1.16
	Private consumption			
Δ income taxes	-0.0068	-0.300	0.044	1.37
Δ transfers	0.149	2.32	0.142	1.77
Δ social security	-0.113	-1.92	-0.02	-0.40

Note : The data in bold figure out the effects that are significant.

In light of the recent crisis, using the expenditure as the main instrument of the budgetary policy in order to cope with the drop of the real GDP and the employment rate was probably a better choice than a policy favoring recovery through fiscal cuts. Our results show evidence in favor of such a strategy. Though tax cuts reduce the risk of a depression by raising the real GDP, the spending multiplier is larger than the one associated with tax cuts. Further, if we consider fiscal stimulus aimed at consumers and enterprises, a decrease in the direct taxes (corporate taxes or income taxes) is likely not to raise either consumption or private investment in times of crises. For reasons explained earlier, the propensity to spend out of such taxes may be offset by non-Keynesian effects. In the current juncture, transfers to households may help to support consumption which has the greatest contribution to GDP. However, direct subsidies to enterprises, in the current environment may not help due to the sharp fall in demand and the uncertainty facing the firms about how good the economic will be in the future (this explains the negative sign associated with the variable reflecting changes in subsidies).

Our estimates take into account the fact that the reactions of the economy to fiscal measures can be influenced by the growth rate of government debt. Ricardian behaviors are likely to affect the magnitude of the fiscal multiplier only and this explains why we find a higher value for the multiplier of expenditure in comparison with that of fiscal. This means that the budgetary instrument used to influence the economy during crisis and non-crisis is not neutral in terms of the probability of being in either regime or the other. Should a government cut taxes, while increasing its indebtedness, that this strategy would be interpreted as signaling future tax increases, thereby implying a higher likelihood of driving the economy out of an expansion phase. In contrast, in presence of a crisis, raising the expenditure while borrowing more might be interpreted as a way of increasing a Government room for manoeuvre, which will stimulate the economy in escaping from a recession. Extrapolating these results, it seems that the increase in public spending corresponding to a large part of the stimulus plans in 2009 (during a recession period) was likely to give way to a rise in GDP growth. On the contrary, the use of the tax cuts would not have produced significant results on GDP growth.

Beyond the crisis, sustainability concerns will be essential for the French government. This could be achieved as follows. The French government could increase the scope for automatic stabilizers and therefore make the discretionary spending measures reversible. Regarding our results, such a strategy could allow to reduce deficits without negative effects on the economy since in times of non-crisis, the multiplier associated with changes in the differences between changes in government spending and the growth rate of potential output is not statistically significant. Considerations could also be given to higher taxes since they do not seem to be a threat for a decrease in the real GDP in the short term (we found no significant effects associated with government revenues in non-crisis time). But, the government would need to target the tax increases. This consideration is important given the ongoing debate on the “fiscal shield”. On the one hand, higher direct taxes on firms could force them to cut investment and employment, as reflected by the negative coefficients associated with corporate taxes and the unit labor costs in the non-crisis regime. On the other hand considering increase in direct taxes on consumers would probably not shift their spending.

6.- Conclusion

It should be reminded that the only empirical models likely to give directly policy implications are structural, such as macro-econometric models or simulation models like DSGE type (but they are accused of ideas based on a priori). The models based on reduced forms (which include all VAR models) are simply intended to give a certain number of facts on which we can base the formulation of economic policy. From this point of view, our study based on TVPMS models allows to highlight several interesting points. The analysis of the role of fiscal variables on some major macroeconomic variables through a TVPMS model clearly shows asymmetry in the effects of fiscal variables depending upon whether one is in periods of crisis or good times. These nonlinearities are both

frequent (as they exist on all behaviors analyzed: GDP, private consumption, business investment and private employment) and significant.

In particular, if one considers the aggregate GDP, public expenditure has a stronger impact during crisis and the expenditure multiplier is greater than the tax multiplier. The consequence is that, during a crisis, a stimulus plan expenditure-oriented might be more efficient than a recovery plan based on measures of tax relief. The effect of tax-oriented measures is significant when the endogenous variables are private investment and employment.

When households are sensitive to the unemployment situation, tax cuts do not affect increase consumption spending, while transfers are playing a significant role. In terms of economic policy, assuming for example that the government's exit strategy consists in stimulating private consumption, it has to choose between two instruments: on the one hand, an increase in transfer expenditure financed by borrowing and, on the other hand lower taxes paid by households.

On the firms side, our results show that direct taxes changes induce a (stimulus) effect in the investment rate only during non-crisis periods. A rise in subsidies has a negative influence during crises, as firms reduce their production capacity.

Increased public spending appears to have a strong multiplier effect at the aggregate level, but with crowding-out effects observed on private investment in non-crisis times. Finally, the estimates suggest that employment policies should be asymmetric: fiscal measures aiming at reducing unit labor costs could be efficient in good times, while an increase in public employment is preferable during crisis.

References

- Claeys, P. (2008). "Rules and their effects on fiscal policy", *Swedish Economic Policy Review*, 15(1), 7-48.
- Cogan, J., Cwik, T., Taylor, J., Wieland, V. (2006). "New Keynesian versus old Keynesian government spending multiplier", *Journal of Economic Dynamics and Control*, 34(3), 281-295.
- Davig, T. (2004), "Regime-switching debt and taxation", *Journal of Monetary Economics*, 51, 837-859.
- Diebold F.X., Lee J-H., Weinbach G.C. (1994). "Regime switching with Time-Varying Transitions Probabilities, in Hargreaves, C.P. (eds) : *Nonstationary Time Series Analysis and Cointegration*, pp. 283-302.
- Dotsey, M. (1994). « Some unpleasant supply-side arithmetic", *Journal of Monetary Economics*, 33, 507-524.

- Dotsey, M., Mao, C. (1997). "The effects of fiscal policy in a neoclassical growth model", Federal Reserve Bank of Richmond Working Paper 97-8.
- Favero, C., Monacelli, T. (2005). "Fiscal policy rules and regime (in)stability: evidence from the US", IGER Working Paper Series 282.
- Filardo A.J. (1994). "Business-cycle Phases and their Transitional Dynamics", *Journal of Business and Economic Statistics* 12(3), pp. 299-308.
- Filardo, A.J., Gordon, S.F. (1993). "Business Cycle Durations," Papers 9328, *Laval - Recherche en Politique Economique*.
- Garcia R. (1998) "Asymptotic Null Distribution of the Likelihood Ratio Test in Markov Switching Models", *International Economic Review*, 39(3), August, 763-788.
- Giavazzi, F., Jappelli T., Pagano M. (2000). "Searching for Non-Linear Effects of Fiscal Policy: Evidence from Industrial and Developing Countries", *European Economic Review* 44 (7), pp. 1259-1290.
- Giavazzi F., Jappelli T., Pagano M., Benedetti M. (2005). "Searching for Non-Monotonic Effects of Fiscal Policy: New Evidence," *CSEF Working Papers 142*, Centre for Studies in Economics and Finance (CSEF), University of Naples, Italy
- Goldfeld, S.M. & Quandt R.M. (1973). "A Markov Model for Switching Regressions", *Journal of Econometrics*, 1, 3-16.
- Hamilton, J.D. (1989). "A New Approach to the Economic Analysis of Nonstationary Time Series and the Business Cycle", *Econometrica*, 57, 357-384.
- Hamilton, J.D. (1991). "A Quasi-Bayesian Approach to Estimating Parameters for Mixtures of Normal Distributions", *Journal of Business and Economic Statistics*, 9, 27-39.
- Hansen, B.E. (1992). "The Likelihood Ratio Test Under Non-Standard Conditions: Testing the Markov Trend Model of GNP", *Journal of Applied Econometrics*, 7, \$61-\$82.
- Hinish M.J., Patterson D.M. (1989) "Evidence of nonlinearity in the trade-by-trade stock market return generating process" in Barnett W.A., Geweke J. and Shell K. (eds) *Economic complexity: chaos, sunspots, bubbles and nonlinearity*, international symposium in Economic theory and Econometrics, chapter 16, 383-409, Cambridge University Press.
- Kiefer N.M. (1978). "Discrete parameter variation: Efficient estimation of a switching regression model", *Econometrica* 46, pp. 427-434.
- Kim C.-J., Piger J., Startz R. (2008). "Estimation of Markov regime-switching regression models with endogenous switching", *Journal of Econometrics* 143 (2) (2008), pp. 263-273.

- Minea, A., Villieu, P. (2008), "Un réexamen de la relation non-linéaire entre déficits budgétaires et croissance économique", *Revue Economique*, 59(3), 561-570.
- Perotti R. (1999). "Fiscal Crisis and Aggregate Demand: Can High Public Debt Reverse the Effects of Fiscal Policy?", *Journal of Public Economics* 65, pp. 147-162.
- Ruge-Murcia, T. (1995). "Credibility and changes in regime", *Journal of Political Economy*, 103, 176-208.
- Sutherland, A. (1997). "Fiscal crises and aggregate demand: can high public debt reverse the effects of fiscal policy?", *Journal of Public Economics*, 65, 147-161.
- Tagkalakis, A. (2008). "The effects of fiscal policy on consumption in recessions and expansions," *Journal of Public Economics, Elsevier*, vol. 92(5-6), pages 1486-1508, June.
- Thams, A. (2006). "Fiscal effects in the European Union", SFB Discussion Paper n°649.
- Trabandt, M., Uhlig, H. (2009). "How far are we from the slippery slope?", NBER Working Paper n°15343.
- Tsay R.S. (1996), "Nonlinearity Tests for Time Series", *Biometrika*, vol 73, 461-466.
- Uhlig, H. (2009). "Some fiscal calculus", University of Chicago, mimeo.

Table 1a. – Real GDP- TVPMS model for France – 1979:01-2009:04
 Budgetary variable : Δ government spending

Explanatory variable	Coefficient	T-ratio	p-value
Intercept (regime 1)	0.005	2.26	0.023
Intercept (regime 2)	-0.013	-5.13	0.0
AR(1) coefficient (regime 1)	0.335	3.43	0.0
AR(1) coefficient (regime 2)	-0.196	-0.99	0.322
Residual standard error (regime 1)	0.005	14.60	0.0
Residual standard error (regime 2)	0.003	2.01	0.04
Δ government spending(t-2) (regime 1)	0.248	2.753	0.005
Δ government spending(t-2) (regime 2)	0.370	3.947	0.0
Δ degree of openness (t-1)	0.047	1.828	0.067
Real interest rate(t-1)	-0.0008	-1.019	0.308
Transition variable : Δ debt(t-1) (smoothed)			
A1	8.77	2.59	0.009
A2	-1.35	-0.25	0.799
B1	-255.18	-1.847	0.064
B2	67.44	0.322	0.746
Likelihood ratio test for TVPMS (null hypothesis: constant probabilities)			
Chi-squared(2) : 8.834 with significance level 0.01206			
Tests on residuals			
Ljung-Box statistics (autocorrelation of order k) : LB(k)			
LB(1): 1.134 significance level: 0.286			
LB(2): 1.552 significance level : 0.46			
LB(3): 1.568 significance level: 0.666			
Linearity tests			
Hinich bispectral test (statistics and p-value) : -3.285 0.99			
Tsay test (statistics and p-value) : 2.917 0.001			

Table 1b. – Real GDP- TVPMS model for France – 1979:01-2009:04

Budgetary variable : $\Delta \text{spendgap} = \Delta \text{government spending} - \Delta \text{potential output}$

Explanatory variable	Coefficient	T-ratio	p-value
Intercept (regime 1)	0.004	2.424	0.015
Intercept (regime 2)	-0.009	-3.823	0.0001
AR(1) coefficient (regime 1)	0.148	1.422	0.155
AR(1) coefficient (regime 2)	-0.177	-0.654	0.512
Residual standard error (regime 1)	0.005	14.16	0.0
Residual standard error (regime 2)	0.004	3.45	0.0
$\Delta \text{spendgap}(t-5)$ (regime 1)	0.05	1.01	0.31
$\Delta \text{spendgap}(t-5)$ (regime 2)	0.296	2.45	0.014
Δ degree of openness (t-1)	0.073	3.025	0.002
Real interest rate(t-1)	0.0005	0.570	0.568
Transition variable : $\Delta \text{debt}(t-2)$ (smoothed)			
A1	8.62	2.62	0.008
A2	0.316	0.068	0.945
B1	-270.62	-1.843	0.065
B2	26.23	0.134	0.893
Likelihood ratio test for TVPMS (null hypothesis: constant probabilities)			
Chi-squared(2) : 5.331 with significance level 0.0695			

Tests on residuals

Ljung-Box statistics (autocorrelation of order k) : LB(k)

LB(1): 1.474 significance level: 0.224

LB(2): 2.492 significance level : 0.287

LB(3): 4.116 significance level: 0.249

Linearity tests

Hinich bispectral test (statistics and p-value) : 2.429 0.0075

Tsay test (statistics and p-value) : 0.983 0.476

Table 1c. – Real GDP- TVPMS model for France – 1979:01-2009:04
 Budgetary variable : $\Delta(\text{Government revenues}/\text{GDP})$

Explanatory variable	Coefficient	T-ratio	p-value
Intercept (regime 1)	-0.010	-3.272	0.001
Intercept (regime 2)	0.006	3.345	0.0008
AR(1) coefficient (regime 1)	0.0209	0.069	0.944
AR(1) coefficient (regime 2)	0.186	2.11	0.034
Residual standard error	0.005	14.957	0.0
Δ government revenues/GDP(t-1) (regime 1)	-0.257	-2.19	0.027
Δ government revenues/GDP(t-1) (regime 2)	-0.044	-1.032	0.302
Δ degree of openness (t-1)	0.058	2.293	0.021
Real interest rate(t-1)	-0.0008	-0.922	0.356
Transition variable : Δ debt ratio(t-1)			
A1	1.019	0.759	0.44
A2	5.743	3.798	0.0001
B1	-24.47	-0.777	0.436
B2	-111.11	-2.511	0.012
Likelihood ratio test for TVPMS (null hypothesis: constant probabilities)			
Chi-squared(2) : 6.278 with significance level 0.043			
Tests on residuals			
Ljung-Box statistics (autocorrelation of order k) : LB(k)			
LB(1): 1.093 significance level: 0.295			
LB(2): 3.001 significance level : 0.222			
LB(3): 4.35 significance level: 0.226			
Linearity tests			
Hinich bispectral test (statistics and p-value) : -0.343 0.634			
Tsay test (statistics and p-value) : 2.04 0.021			

Table 2. – Real private consumption- TVPMS model for France – 1970:01-2009:04

Explanatory variable	Coefficient	T-ratio	p-value
Intercept (regime 1)	0.00031	0.348	0.727
Intercept (regime 2)	0.006	5.986	0.0
AR(1) coefficient (regime 1)	0.027	0.164	0.869
AR(1) coefficient (regime 2)	-0.243	-2.08	0.037
Residual standard error	0.003	10.41	0.0
Δ income taxes(t) (regime 1)	-0.0068	-0.300	0.763
Δ income taxes(t) (regime 2)	0.044	1.369	0.170
Δ transfers(t-1) (regime 1)	0.149	2.319	0.02
Δ transfers(t-1) (regime 2)	0.142	1.768	0.076
Δ social security(t) (regime 1)	-0.113	-1.919	0.054
Δ social security(t) (regime 2)	-0.02	-0.401	0.688
Δ real disposable income	0.139	2.158	0.03
Transition variable : unemployment rate (smoothed)			
A1	-0.234	-0.354	0.723
A2	1.319	2.02	0.043
B1	163.83	2.793	0.0052
B2	-22.97	-0.543	0.586
Likelihood ratio test for TVPMS (null hypothesis: constant probabilities)			
Chi-squared(2) : 8.238 with significance level 0.0162			
Tests on residuals			
Ljung-Box statistics (autocorrelation of order k) : LB(k)			
LB(1): 0.244 significance level: 0.62			
LB(2): 1.695 significance level : 0.428			
LB(3): 1.805 significance level: 0.613			
Linearity tests			
Hinich bispectral test (statistics and p-value) : -1.968 0.975			
Tsay test (statistics and p-value) : 2.079 0.019			

Table 3. – Business investment- TVPMS model for France – 1970:01-2009:04

Explanatory variable	Coefficient	T-ratio	p-value
Intercept (regime 1)	0.008	2.55	0.01
Intercept (regime 2)	-0.003	-0.77	0.438
AR(1) coefficient (regime 1)	0.012	0.11	0.905
AR(1) coefficient (regime 2)	0.276	2.57	0.01
Residual standard error	0.01	14.57	0.0
Δ corporate taxes(t-3) (regime 1)	-0.08	-2.21	0.027
Δ corporate taxes(t-3) (regime 2)	0.022	0.76	0.442
Δ subsidies(t-2) (regime 1)	0.048	1.27	0.201
Δ subsidies(t-2) (regime 2)	-0.17	-3.04	0.0023
Δgovernment spending(t-3) (regime 1)	-0.394	-2.422	0.015
Δgovernment spending(t-3) (regime 2)	-0.357	-1.16	0.244
Δ real GDP(t-2)	0.430	1.928	0.053
Δ real GDP(t-3)	1.13	5.25	0.0
Real long-run interest rate (t-2)	-0.001	-3.38	0.0
Transition variable : output-gap			
A1	2.07	2.73	0.006
A2	2.59	3.80	0.0
B1	-1.063	-1.52	0.127
B2	1.036	2.58	0.009
Likelihood ratio test for TVPMS (null hypothesis: constant probabilities)			
Chi-squared(2) : 9.524 with significance level 0.0085			
Tests on residuals			
Ljung-Box statistics (autocorrelation of order k) : LB(k)			
LB(1): 0.212 significance level: 0.644			
LB(2): 5.532 significance level : 0.063			
LB(3): 5.716 significance level: 0.126			
Linearity tests			
Hinich bispectral test (statistics and p-value) : -3.313 0.99			
Tsay test (statistics and p-value) : 2.624 0.0029			

Table 4. – Private employment- TVPMS model for France – 1970:01-2009:04

Explanatory variable	Coefficient	T-ratio	p-value
Intercept (regime 1)	-0.002	-4.91	0.0
Intercept (regime 2)	-0.0008	-2.83	0.004
AR(1) coefficient (regime 1)	0.579	13.76	0.0
AR(1) coefficient (regime 2)	1.144	11.83	0.0
Residual standard error (regime 1)	0.00078	11.90	0.0
Residual standard error (regime2)	0.00073	7.57	0.0
Δ unit labor cost(t-3) (regime 1)	0.033	4.19	0.0
Δ unit labor cost(t-3) (regime 2)	-0.122	-8.38	0.0
Δ public investment (t-3) (regime 1)	0.016	3.25	0.001
Δ public investment (t-3) (regime 2)	-0.028	-3.64	0.0002
Δ real GDP(t-1)	0.10	6.68	0.0
Transition variable : output-gap			
A1	0.916	2.00	0.044
A2	-0.741	-1.40	0.161
B1	-0.829	-2.134	0.032
B2	0.132	0.523	0.60
Likelihood ratio test for TVPMS (null hypothesis: constant probabilities)			
Chi-squared(2) : 5.766 with significance level 0.0559			
Tests on residuals			
Ljung-Box statistics (autocorrelation of order k) : LB(k)			
LB(1): 2.366 significance level: 0.123			
LB(2): 2.416 significance level : 0.298			
LB(3): 3.907 significance level: 0.27			
Linearity tests			
Hinich bispectral test (statistics and p-value) : 1.621 0.0525			
Tsay test (statistics and p-value) : 2.053 0.0182			

Figure 1. Posterior smoothed probability of regime 2 (huge falls in real GDP)
Budgetary variable : government spending

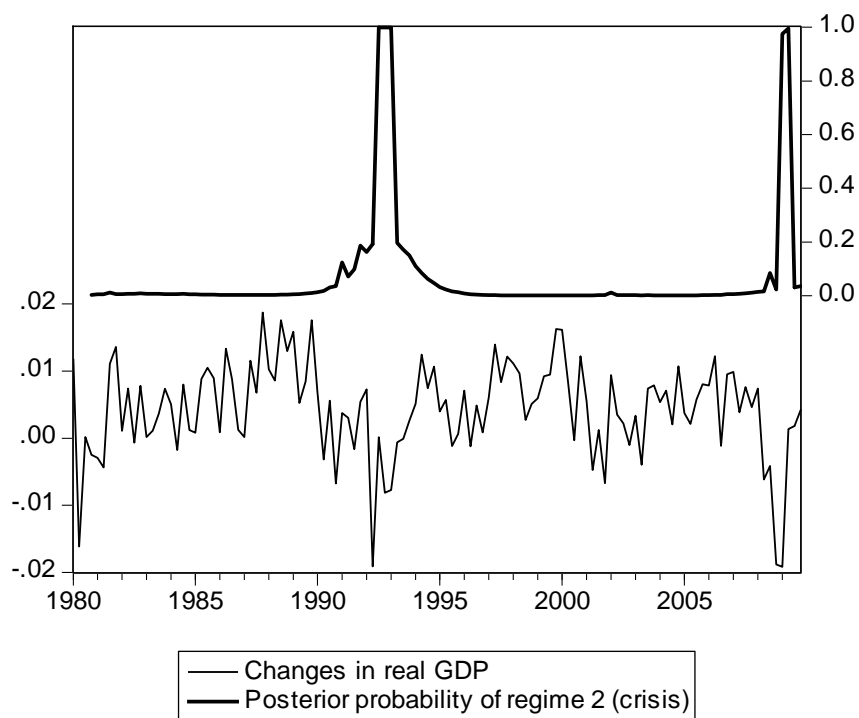


Figure 2. Posterior smoothed probability of regime 2 (huge falls in real GDP)
Budgetary variable : government spending-potential growth

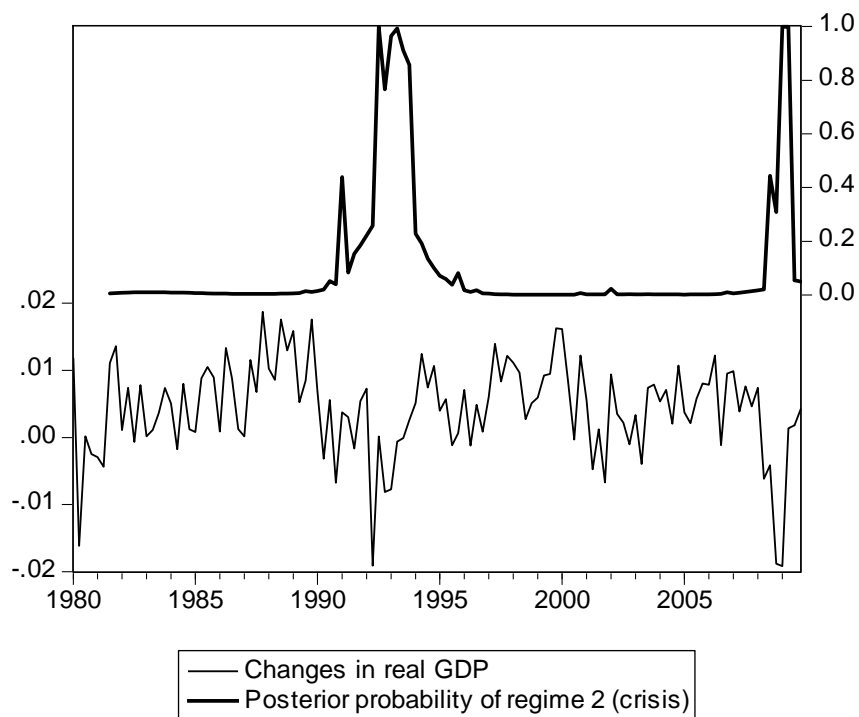


Figure 3. Posterior smoothed probability of regime 1 (huge falls in real GDP)
 Budgetary variable : government revenues/GDP

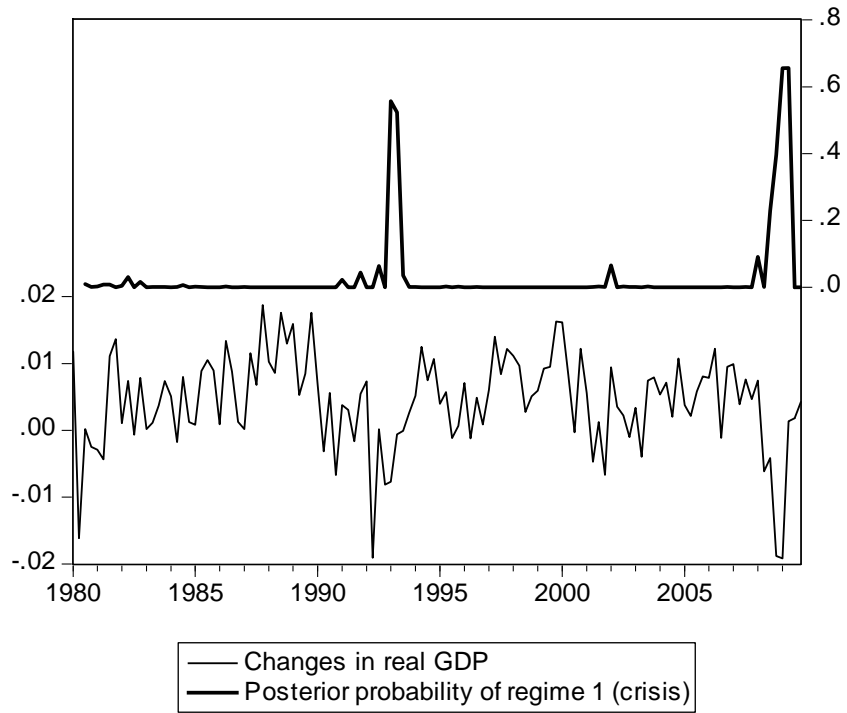


Figure 4a. Posterior smoothed probability of regime 1 (huge falls in private consumption)

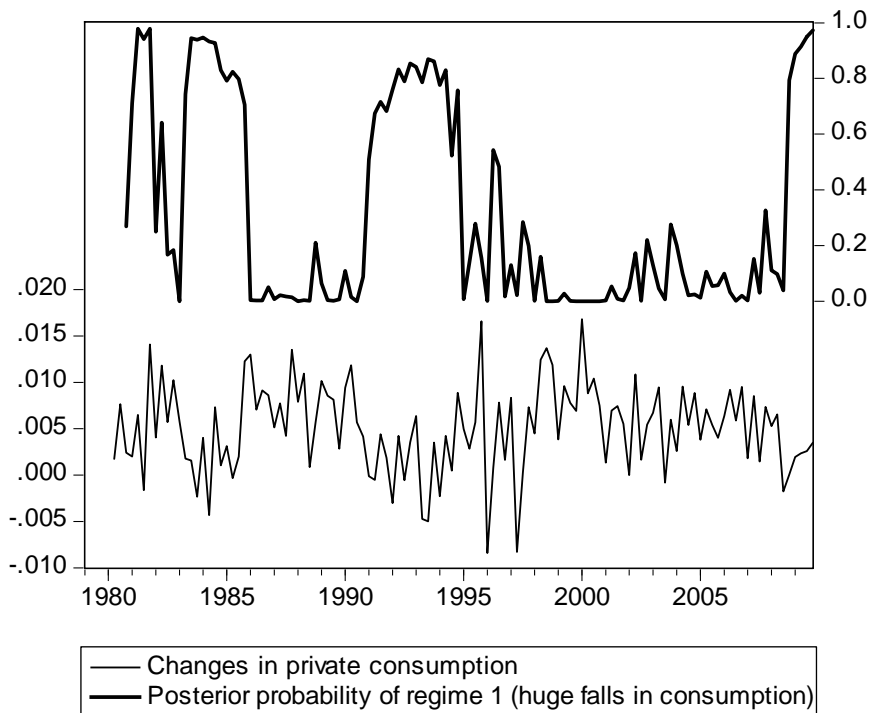


Figure 4b. Posterior smoothed probability of regime 2 (other periods of than huge falls in private consumption)

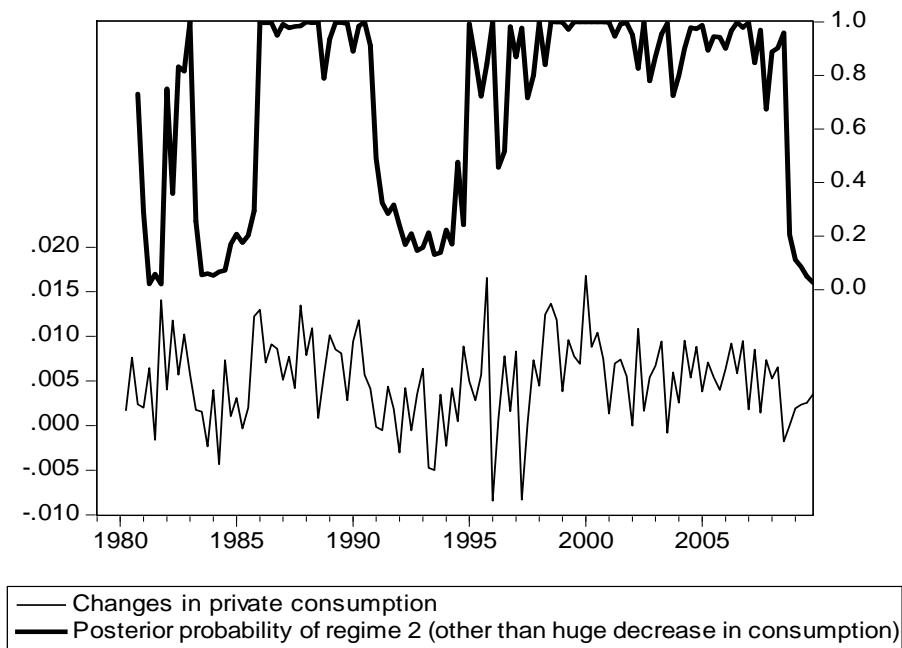


Figure 5a. Posterior smoothed probability of regime 2 (prolonged decreased in business investment) and changes in business investment

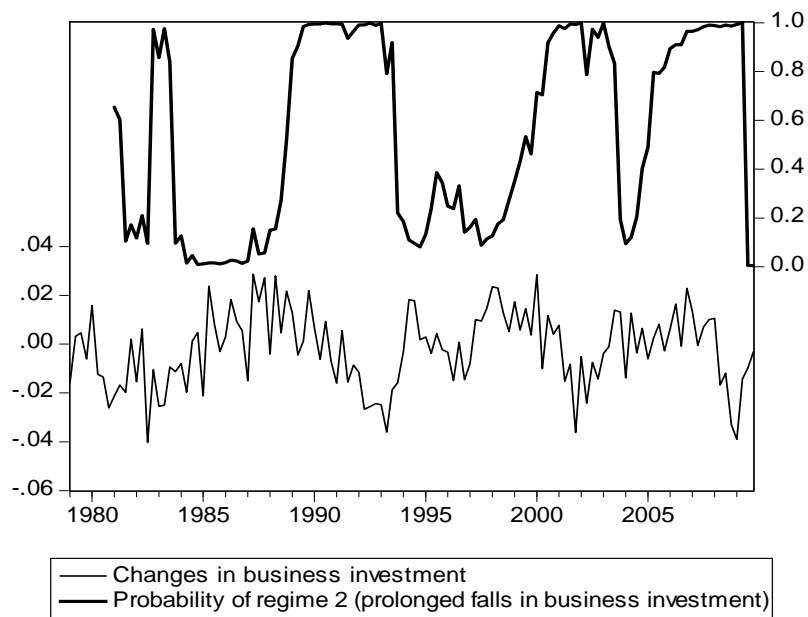


Figure 5b. Posterior smoothed probability of regime 1 (sustained increase in business investment) and changes in business investment

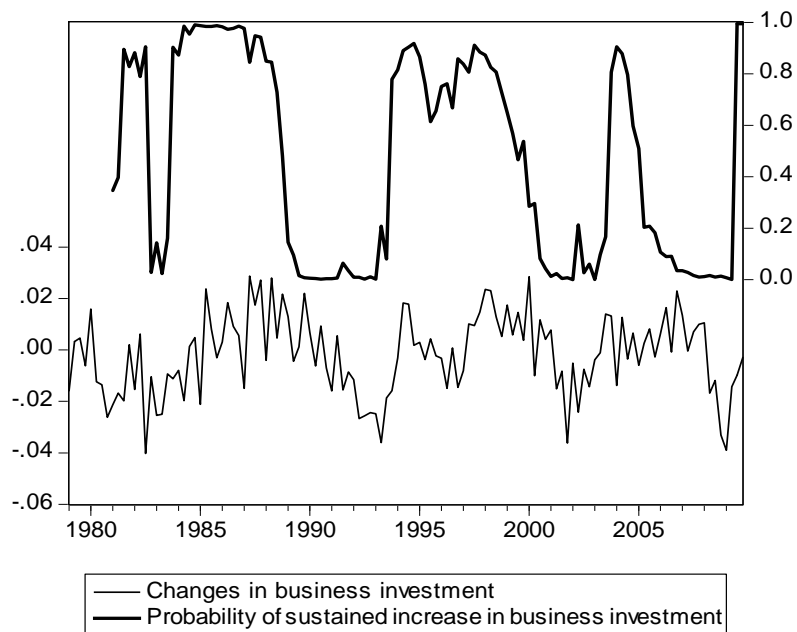


Figure 6a. Private employment in logarithm. Raw series and estimated trend using HP

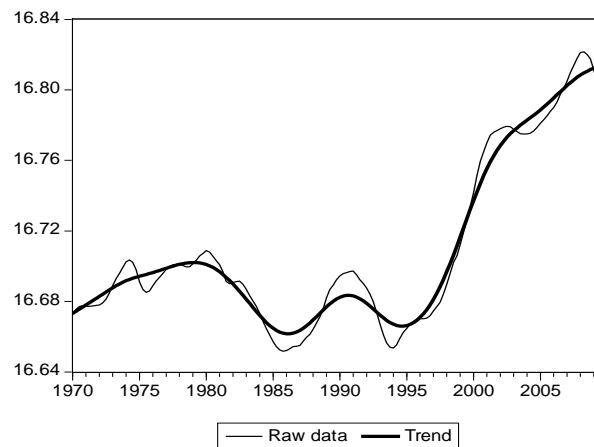


Figure 6b. Private employment in logarithm. –Detrended series

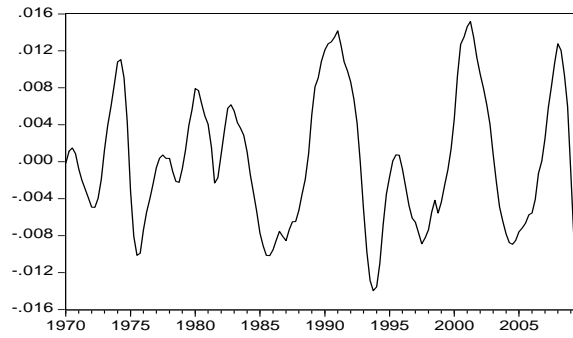
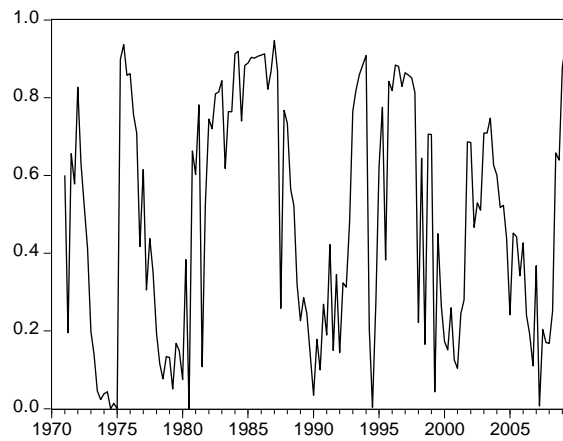


Figure 6c. Steepness of employment variations and output-gap



Figure 6d. Posterior smoothed probability of regime 1 (times of crisis)



Appendix. – Data description and definition of the variables

Data are at a quarterly frequency and are taken from the OECD database. Public finance data are initially available at an annual frequency from the OECD database and were then interpolated to obtain quarterly data. All data span from 1978:1 to 2009:2. The codes are those of the OECD database.

1.- Real GDP equation

Endogenous variable

- $\Delta \log(\text{real private GDP}(t))$ defined as follows
$$\Delta \log \left\{ \frac{\text{FRA_GDP}(t) - \text{FRA_CGAA}(t)}{\text{FRA_PGDP}(t)} \right\}$$
where
FRA_GDP : nominal GDP
FRA_CGAA : public sector consumption
FRA_PGDP : GDP deflator

Control variables

- $\Delta \log(\text{openness}(t))$ defined as follows
$$\Delta \log \left\{ \frac{\text{FRA_XGS}(t) + \text{FRA_MGS}(t)}{\text{FRA_GDP}(t) + \text{FRA_MGS}(t)} \right\}$$
where
FRA_GDP : nominal GDP
FRA_XGS : exports of goods and services, local currency
FRA_MGS : imports of goods and services, local currency
- $\log(1 + \text{real interest rate}(t))$ defined as follows
$$\log(1 + \text{FRA_IRS}(t) - \Delta \text{FRA_CPI}(t) / \text{FRACPI}(t-1))$$
where
FRA_CPI : consumer price index
FRA_IRS : short-term interest rate, nominal

Public finance variables

- $\Delta \log(\text{real government spending}(t))$ defined as follows
$$\Delta \log(\text{FRA_YPGT}(t) / \text{FRA_PGDP}(t))$$
where
FRA_YPGT : public sector total expenditure
FRA_PGDP : GDP deflator
- $\Delta \log(\text{real government revenues}(t))$ defined as
$$\Delta \log(\text{FRA_YRGT}(t) / \text{FRA_PGDP}(t))$$
where
FRA_YRGT : public sector total revenues
FRA_PGDP : GDP deflator
- $\log(\text{real government spending}(t)) - \log(\text{potential output}(t))$ defined
Potential output is taken from OECD series.

Transition variables

- Changes in public debt and changes in public debt ratio, defined as the trend component of $\Delta\log(\text{public debt}(t))$ or $\Delta\log(\text{public debt ratio}(t))$ using an HP filter

2.- Real private consumption

Endogenous variable

- $\Delta\log(\text{real private consumption}(t))$ defined as follows
 $\Delta\log\{FRA_CPV(t)\}$
where FRA_CPV : private consumption, constant price

Control variables

- $\Delta\log(\text{real disposable income}(t))$ defined as
 $\Delta\log\{FRA_YDRH(t)\}$
where FRA_YDRH : households' real disposable income

Public finance variables

- $\Delta\log(\text{real taxes}(t))$ defined as
 $\Delta\log\{FRA_TYHA(t)/FRA_CPI(t)\}$
where
FRA_TYHA : direct taxes paid by households (cyclically adjusted)
FRA_CPI : consumer price index
- $\Delta\log(\text{real transfers received by households})$ defined as
 $\Delta\log\{TRRH(t)/FRA_CPI(t)\}$
where
FRA_TRRH : net transfers received by households, nominal
FRA_CPI : consumer price index
- $\Delta\log(\text{households' contribution to social security}(t))$ defined as
 $\Delta\log\{FRA_TRSSH(t)/FRA_CPI(t)\}$
where
FRA_TRSSH : households' contribution to social security, nominal
FRA_CPI : consumer price index

Transition variable

- Changes in unemployment rate (smoothed) defined as the trend component of $\Delta\log(\text{unemployment rate}(t))$ using an HP filter.

3.- Business investment

Endogenous variable

- $\Delta\log(\text{firms' investment rate}(t))$ defined as follows
 $\Delta\log\{FRA_IBV(t)/FRA_KTVS(t)\}$
where
FRA_IBV : private non residential fixed capital, constant prices
FRA_KTVS : smoothed capital stock

Control variables

- $\Delta \log(\text{real GDP}(t))$. See the definition above
- $\log(1 + \text{real interest rate}(t))$. See the definition above

Public finance variables

- $\Delta \log(\text{real taxes}(t))$ defined as:
 $\Delta \log(\text{FRA_TYB}(t)/\text{FRA_PGDP}(t))$
where
FRA_TYB : direct taxes paid by the enterprises
FRA_PGDP : GDP deflator
- $\Delta \log(\text{subsidies}(t) \text{ in real terms})$ defined as
 $\Delta \log(\text{FRA_TSUB}(t)/\text{FRA_PGDP}(t))$
where
FRA_TSUB : subsidies, nominal
FRA_PGDP : GDP deflator
- $\Delta \log(\text{real government spending}(t))$. See the definition above.

Transition variable

- output-gap (computed by the OECD)

4.-Private employment

Endogenous variable

- Steepness of changes in private employment defined as follows:
Step 1 : we compute $\log(\text{FRA_ETB}(t))$
where FRA_ETB : total employment, private sector
Step 2 : we apply an HP filter to $\log(\text{FRA_ETB}(t))$ and obtain the cyclical component named LETB_CYCLE(t)
Step 3 : we compute $\Delta \text{LETB_CYCLE}(t)$

Control variable

- $\Delta \log(\text{real GDP}(t))$ defined as
 $\Delta \log(\text{FRA_GDP}(t)/\text{FRA_PGDP}(t))$
where
FRA_GDP : GDP, nominal
FRA_PGDP : GDP deflator

Public finance variables

- $\Delta \log(\text{unit labor cost}(t))$
- Changes in public investment ratio defined as follows:
 $\Delta \log(\text{FRA_IGV}(t)/\text{FRA_GDP}(t))$
where
FRA_IGV : public non residential fixed investment
FRA_GDP : GDP, nominal

Transition variable : output-gap (computed by OECD).

Documents de Travail

271. G. Cette, Y. Kocoglu, and J. Mairesse, "Productivity Growth and Levels in France, Japan, the United Kingdom and the United States in the Twentieth Century," January 2010
272. E. Lavallée and V. Vicard, "National borders matter...where one draws the lines too," January 2010
273. C. Loupias and P. Sevestre, "Costs, demand, and producer price changes," January 2010
274. O. de Bandt, K. Barhoumi and C. Bruneau, "The International Transmission of House Price Shocks," January 2010
275. L. Ferrara and S. J. Koopmany, "Common business and housing market cycles in the Euro area from a multivariate decomposition," January 2010
276. V. Chauvin and O. Damette, "Wealth effects: the French case," January 2010
277. H. Dellas, B. Diba and O. Loisel, "Financial Shocks and Optimal Policy," March 2010
278. S. Avouyi-Dovi and J. Idier, "Central bank liquidity and market liquidity: the role of collateral provision on the French government debt securities market," January 2010
279. V. Borge, J. Idier and G. Le Fol, "Liquidity problems in the FX liquid market: Ask for the "BIL"," March 2010
280. D. Durant et L. Frey, "Une Première comparaison des droits à pension des ménages français et américains," Avril 2010
281. G. Bertola, A. Dabusinkas, M. Hoeberichts, M. Izquierdo, C. Kwapil, J. Montornès and D. Radowski, "Price, Wage and Employment Response to Shocks: Evidence from the WDN Survey," May 2010
282. J. Montornès and J-B. Sauner-Leroy, "Wage-setting Behavior in France: Additional Evidence from an Ad-hoc Survey," May 2010
283. R. Bourlès, G. Cette, J. Lopez, J. Mairesse and G. Nicoletti, "Do product market regulations in upstream sectors curb productivity growth? Panel data evidence for OECD countries," June 2010
284. Ph. Askenazy, Th. Breda and D. Irac, "Innovation and Advertising: Theory and Evidence," May 2010
285. M. Lemoine and C. Mougin, "The Growth-Volatility Relationship: New Evidence Based on Stochastic Volatility in Mean Models," July 2010
286. C. Bouthevillain and G. Dufrénot, "Are the effects of fiscal changes different in times of crisis and non-crisis? The French Case," July 2010

Pour accéder à la liste complète des Documents de Travail publiés par la Banque de France veuillez consulter le site : http://www.banque-france.fr/fr/publications/documents_de_travail/documents_de_travail_10.htm

For a complete list of Working Papers published by the Banque de France, please visit the website: http://www.banque-france.fr/fr/publications/documents_de_travail/documents_de_travail_10.htm

Pour tous commentaires ou demandes sur les Documents de Travail, contacter la bibliothèque de la Direction Générale des Études et des Relations Internationales à l'adresse suivante :

For any comment or enquiries on the Working Papers, contact the library of the Directorate General Economics and International Relations at the following address :

BANQUE DE FRANCE
49- 1404 Labolog
75049 Paris Cedex 01
tél : 0033 (0)1 42 92 49 55 ou 62 65 ou 48 90 ou 69 81
email : thierry.demoulin@banque-france.fr
jeannine.agoutin@banque-france.fr
<mailto:michaelbrassart@banque-france.fr>
veronique.jan-antuoro@banque-france.fr
[nathalie.bataille-salle@banque-france.f](mailto:nathalie.bataille-salle@banque-france.fr)